

1. GMM : Gaussian Mixture Model

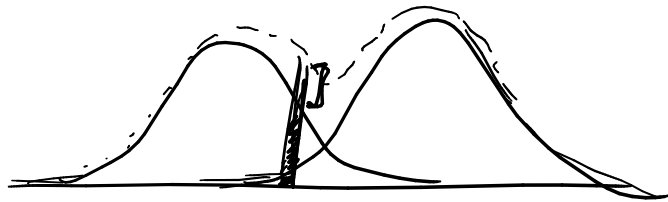
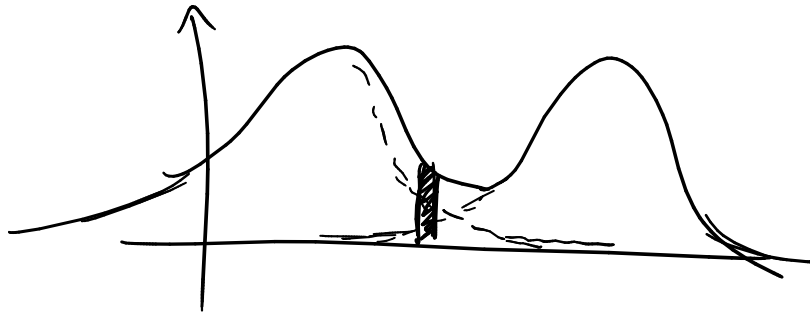
GMM-2.

1.  $Y_1 \sim N(\mu_1, \sigma_1^2)$

2.  $Y_2 \sim N(\mu_2, \sigma_2^2)$

3.  $Y = (1-\Delta)Y_1 + \Delta Y_2$

where  $\Delta \sim \text{Bern}\left(\frac{1}{2}\right)$



$$P(Y=y) = (1-\pi) \underbrace{N(y; \mu_1, \sigma_1^2)} + \pi N(y; \mu_2, \sigma_2^2)$$

$$\frac{1}{\sqrt{2\pi} \sigma_1} \exp\left(-\frac{(y-\mu_1)^2}{2\sigma_1^2}\right)$$

$$\log P(Y=y) = \log\left((1-\pi) \exp(\quad) + \pi \exp(\quad)\right)$$

$$\begin{pmatrix} y_1, \delta_1 \\ y_2, \delta_2 \\ \vdots \\ y_N, \delta_N \end{pmatrix}$$

1st

$Y, \Delta$

$$\Delta \in \{0, 1\}$$

$$\Delta \sim B(\underline{\pi})$$

$$\begin{pmatrix} y_1, \delta_1 \\ \vdots \\ y_N, \delta_N \end{pmatrix}$$

2nd

$y_i, \delta_i$

$$LL \left( \begin{pmatrix} \pi \\ \mu_1 \\ \sigma_1 \\ \mu_2 \\ \sigma_2 \end{pmatrix} \right) = \sum_{i=1}^N \left[ \begin{array}{l} \uparrow \\ (1 - \delta_i) \left( \log N(y_i; \mu_1, \sigma_1^2) + \log(1 - \pi) \right) \\ + \delta_i \left( \log N(y_i; \mu_2, \sigma_2^2) + \log \pi \right) \end{array} \right]$$

$\underbrace{\quad}_{\theta \in \mathbb{R}^5}$

max  
 $\pi$

$$\sum_{i=1}^N (1-\delta_i) \log(1-\pi) + \delta_i \log \pi$$

$M$

$$\sum_{i=1}^N (1-\delta_i) \frac{-1}{(1-\pi)} + \sum_{i=1}^M \delta_i \frac{1}{\pi} = 0$$

$$\frac{(N-M)}{1-\pi} = \frac{M}{\pi} \Rightarrow \frac{N-M}{M} = \frac{1}{\pi} - 1$$

$$\Rightarrow \frac{N}{M} = \frac{1}{\pi}$$

$$\Rightarrow \pi = \frac{M}{N}$$

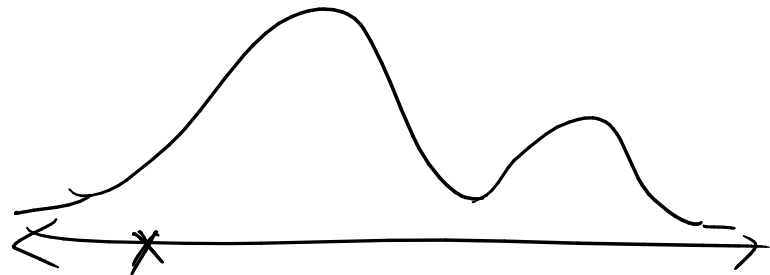
$$\theta = \begin{bmatrix} \pi \\ \mu_1 \\ \sigma_1 \\ \mu_2 \\ \sigma_2 \end{bmatrix}$$

$$\begin{bmatrix} y_1 \\ \vdots \\ y_N \end{bmatrix} = \mathbf{Y}$$

$$E[\Delta_i | \theta, \mathbf{Y}]$$

$$\gamma_i \equiv P(\Delta_i = 1 | \theta, y_i)$$

		<u>weight</u>
$y_1$	1	$\gamma_1$
$y_1$	0	$1 - \gamma_1$
$y_2$	1	$\gamma_2$
$y_2$	0	$1 - \gamma_2$
$\vdots$	$\vdots$	
$\vdots$	$\vdots$	
$y_N$	1	
$y_N$	0	



Why EM works :

	<u><math>Z, Z^m</math></u>	$Y, \Delta$
	$l(\theta; Z)$ $l_0(\theta; Z, Z^m)$	
<u>E.</u>	<u><math>P(Z^m   Z, \theta^{old})</math></u>	$P(\Delta_i = 1   y_i, \theta^{old}) \quad i=1 \dots N$
Weighted LL	$Q(\theta', \theta^{old})$	
<u>M.</u>	$\theta^{New} = \arg \max_{\theta'} \underline{\underline{Q(\theta', \theta^{old})}}$	

$$P(Z; \theta^{\text{new}}) = \frac{P(z^m, Z; \theta^{\text{new}})}{P(z^m | Z, \theta^{\text{new}})} \quad \Bigg| \quad \begin{aligned} &P(z^m | Z; \theta) \\ &= \frac{P(z^m, Z; \theta)}{P(Z; \theta)} \end{aligned}$$

① log

$$\begin{aligned} \ell(Z; \theta^{\text{new}}) &= \ell_0(z^m, Z; \theta^{\text{new}}) \\ &\quad - \ell_1(\theta^{\text{new}}) \end{aligned}$$

②  $E_{P(z^m | Z, \theta^{\text{old}})}$

$$\underbrace{\ell(z; \theta^{\text{new}})} = \underbrace{Q(\theta^{\text{new}}, \theta^{\text{old}})} - \underbrace{R(\theta^{\text{new}}, \theta^{\text{old}})}$$

$$\ell(z, \theta^{\text{old}})$$

$$\underbrace{E_{P(z^m | z, \theta^{\text{old}})} [\ell_1(\theta^{\text{new}})]}$$

$$\ell(\theta^{\text{new}}) - \ell(\theta^{\text{old}}) > 0 \quad \star \star$$

$$\underbrace{Q(\theta^{\text{new}}, \theta^{\text{old}})} - R(\theta^{\text{new}}, \theta^{\text{old}}) > \underbrace{Q(\theta^{\text{old}}, \theta^{\text{old}})} - R(\theta^{\text{old}}, \theta^{\text{old}})$$



$$R(\theta^{\text{new}}, \theta^{\text{old}}) < R(\theta^{\text{old}}, \theta^{\text{old}})$$


$$\underbrace{\sum_{i=1}^k q_i \log p_i} < \underbrace{\sum_{i=1}^k q_i \log q_i}$$

$$\equiv \left. \begin{array}{l} \max_{p_i} \sum_{i=1}^k q_i \log p_i \\ \sum_{i=1}^k p_i = 1 \end{array} \right\}$$

$$\frac{1}{P} \underbrace{\log P(z^m | \theta^{\text{old}})}_q \underbrace{\log P(z^m | \theta^{\text{new}})}_p$$

$$\sum_{i=1}^k q_i \cdot \log p_i$$

$$Obj = \sum_{i=1}^k q_i \log p_i - \lambda \left( \sum_{i=1}^k p_i - 1 \right) \quad p_1, \dots, p_k$$

$$\frac{q_i}{p_i} - \lambda(1) = 0 \Rightarrow \frac{q_i}{p_i} = \lambda \Rightarrow q_i = \lambda p_i$$


$$\sum p_i = 1$$

$$\sum q_i = 1$$

$$\sum q_i = \lambda \sum p_i$$

Sampling

$$\underbrace{P(\beta | \text{data})}_{\text{Sampling}} = \frac{P(\text{data} | \beta) \cdot P(\beta)}{P(\text{data})} = \int_{\beta} P(\text{data} | \beta) \cdot P(\beta)$$

$$Z \sim P_2 \quad z_1, \dots, z_N \sim P_2$$

$$E[Z] \approx \frac{1}{N} \sum_{i=1}^N z_i$$

if we have access to samples from  $U[0,1]$

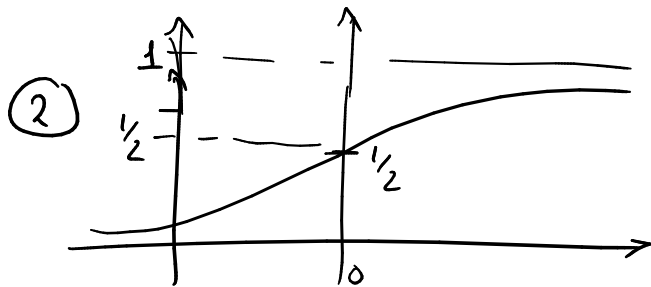
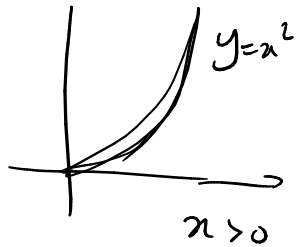
①  $Z \sim B(\frac{1}{2})$

$u \sim U[0,1]$   
if  $(u) \in [0, 0.5]$

$Z = 1$

else

$Z = 0$



$F(x) = P(X \leq x)$

$$U \sim P_U$$

$$u_1, \dots, u_N$$

$$U \in \mathbb{R}^3$$

Assume:  $P(U_i | U_1, \dots, U_{i-1}, U_{i+1}, \dots, U_p)$

Gibbs Sampling.

$$P(U_1 | U_2 U_3)$$

$$P(U_2 | U_1 U_3)$$

$$P(U_3 | U_1 U_2)$$

Markov Chain

Monte Carlo

technique.

$U_1$	0	2	3				
$U_2$	1	4	2				
$U_3$	-3	-1	-2				
	$u_1$	$u_2$					
	$i=1$	$i=2$	$i=3$				

Markov Chain

$$P(Z_{\text{next}} | Z_{\text{old}}) \begin{bmatrix} 0.3 & .7 \\ .4 & .6 \end{bmatrix}$$

$$\begin{bmatrix} a & b \end{bmatrix} \begin{bmatrix} .3 & .7 \\ .4 & .6 \end{bmatrix} = \begin{bmatrix} a & b \end{bmatrix}$$